Causality and its Role in Reasoning, Explainability, and Generalizability

Faculty of Mathematics and Computer Science Philipps-Universität Marburg

Lisbon Machine Learning School (LxMLS) July 20th, 2023

Adèle H. Ribeiro

https://adele.github.io/ adele.ribeiro@uni-marburg.de

Recent Breakthroughs in Al

- high-dimensional settings.
- In particular, there are huge progresses in natural processing language, computer vision, and reinforcement learning.

We can learn models that makes predictions extremely well in



Recent Breakthroughs in Al

The New York Times

Driverless Cars 7 to San Jose





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Al system is better than human doctors at predicting breast cancer

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TECHNOLOGY 1 January 2020

By Jessica Hamzelou



Aside from the lidar range-finder unit on its run, unit icles, including this Lexus hybrid, look reasonably conventional. Google



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THE SHIFT

GPT-4 Is Exciting and Scary

Today, the new language model from OpenAI may not seem all that dangerous. But the worst risks are the ones we cannot anticipate.





Current Challenges in Al



A machine with a grasp of cause and effect could learn more like a human, through imagination and regret.

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Why artificial intelligence needs to understand consequences



Judea Pearl – Causality

JUDEA PEARL

MODEED INFERENCE

"Deep learning has instead given us machines with truly impressive abilities but no intelligence. The difference is profound and lies in the absence of a model of reality."

— The Book of Why: The New Science of Cause and Effect

Director of the Cognitive Systems Laboratory at the University of California, Los Angeles.

In 2011, he won the A. M. Turing Award (the highest distinction in computer science and a \$250,000 prize)

"for fundamental contributions to artificial intelligence through the development of a calculus for probabilistic and causal reasoning."

— Association for Computing Machinery (ACM)







Yoshua Bengio – Deep Learning



Professor at the University of Montreal, and the Founder and Scientific Director of Mila – Quebec Al Institute

In 2018, he won the A. M. Turing Award, with Geoffrey Hinton, and Yann LeCun

"for conceptual and engineering breakthroughs that have made deep neural networks a critical component of computing."

— Association for Computing Machinery (ACM)

"Causality is very important for the next steps of progress of machine learning," — interview with IEEE Spectrum, 2020.





Guido W. Imbens & Joshua D. Angrist



Guido W. Imbens

Professor of Applied **Econometrics in Stanford University**



Joshua D. Angrist

Professor of Economics at the Massachusetts Institute of Technology

CAUSAL INFERENCE FOR STATISTICS, SOCIAL, AND BIOMEDICAL **SCIENCES** AN INTRODUCTION **GUIDO W. IMBENS** DONALD B. RUBIN



In 2021, they won the <u>Nobel Prize</u> in Economics (about \$1 million)

"for their methodological contributions to the analysis of causal relationships"

Why causality is so important?

Causality is an essential component in the development of the new generation of Artificial Intelligence methods, allowing important capabilities such as

Explainability: provides a better understanding of the underlying mechanisms, e.g., learning directionality and confounding through causal structure learning.

Reasoning: can determine the effect of *unrealized* interventions rather than just predicting an outcome (i.e., can distinguish between association and causation).

Fairness: captures and disentangles any mechanisms of discrimination that may be present, including direct, indirect-mediated, and indirect-confounded.

Generalizability: allows the transportability of causal effects across different domains.

Data Fusion: provides language and theory to cohesively combine prior knowledge and data from multiple and heterogeneous studies.





Causal Data Science

Goal is to develop language, criteria, and algorithms for:



Causal inference and the data-fusion problem

Elias Bareinboim^{a,b,1} and Judea Pearl^a

^aDepartment of Computer Science, University of California, Los Angeles, CA 90095; and ^bDepartment of Computer Science, Purdue University, West Lafayette, IN 47907

Edited by Richard M. Shiffrin, Indiana University, Bloomington, IN, and approved March 15, 2016 (received for review June 29, 2015)

Data-Fusion: cohesively combining heterogenous datasets, • Causal Inference: inferring the effects of interventions, and **Decision-Making:** making robust and generalizable decisions.

http://causalfusion.net



Causality Theory by Judea Pearl



CAUSA IN STA A Primer

Judea Pearl Madelyn Glymour Nicholas P. Jewell



CAUSAL INFERENCE IN STATISTICS

WILEY





MODELS, REASONING, AND INFERENCE

JUDEA PEARL

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Causality Theory by Judea Pearl



Causality101

Chapter 2.3 - Colliders

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4	Z 170 _r -20		
5	W 90,60		
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8	Z -> Y 0		
9	Y -> W 0		
10			

https://causality101.net/



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Prediction vs Reasoning Statistical Association vs Causation



Predictive Tasks

Yes!

X: Number of firefighters in action Y: Seriousness of fire

 $\rho_{XY} \neq 0 \implies X \text{ is a good predictor of } Y$

$$P(Y = y | X = x) \neq P(Y = y)$$

Observational Probability Distribution

Task: Can I guess how serious/big is the fire by the number of firefighters in action?



Conclusion: The seriousness of fire increases with the number of firefighters.



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Prediction \Rightarrow Decision-Making / Reasoning?



- **Conclusion:** The size of the fire increases with the number of firefighters.
 - In other words, the fewer the firefighters, the smaller the fire.

Should we decrease the number of firefighters to reduce the fire?



Effect of Interventions

X: Number of firefighters in action *Y*: Seriousness of fire

 $\begin{cases} X = f_X(Y, U_X, U_{XY}) \\ Y = f_Y(U_Y, U_{YY}) \end{cases}$

Underlying Model

 $\swarrow Y$ is not a function of XIn other words, Y is not caused by X



Effect of Interventions

X: Number of firefighters in action *Y*: Seriousness of fire

X = x $(X = f_X(Y, U_X, U_{XY}))$ $Y = f_Y(U_Y, U_{YY})$

Underlying Model

Conclusion: we cannot change the size of the fire by changing the number of firefighters.

Y is not a function of XIn other words, Y is not caused by X

Changing X won't change the value of Y

P(Y = y | do(X = x)) = P(Y = y)

Interventional Probability Distribution

The action/intervention on *X*, do(X = x) is independent of *Y*



Structural Causal Model (SCM) EXPLAINABILITY AND THE DATA GENERATING MODEL

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Structural Causal Model (SCM)

 $\langle \mathbf{V}, \mathbf{U}, \mathcal{F}, P(\mathbf{u}) \rangle$, where

- $\mathbf{V} = \{V_1, \dots, V_n\}$: are endogenous variables
- $\mathbf{U} = \{U_1, \dots, U_m\}$: are exogenous variables
- $\mathcal{F} = \{f_1, \dots, f_n\}$: are functions determining **V**, i.e., $v_i \leftarrow f_i(pa_i, u_i)$, where $Pa_i \subseteq \mathbf{V}$ are endogenous causes (parents) of V_i and $U_i \subseteq \mathbf{U}$ are exogenous causes of V_i .
- $P(\mathbf{U})$ is the probability distribution over \mathbf{U} .

Definition: A structural causal model \mathcal{M} (or, data generating model) is a tuple

Assumption: \mathcal{M} is recursive, i.e., there are no feedback (cyclic) mechanisms.



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Effect of Interventions in SCMs

Pre-Interventional/ **Observational SCM**

Observational Distribution $P(\mathbf{V}) \doteq P_{\mathscr{M}}(\mathbf{V})$

Can we **predict** better the value of Y after **observing** que X = x?

 $P(Y = y | X = x) \neq P(Y = y) \implies X$ is **correlated** to Y

Post-Interventional / Interventional SCM

$$\mathscr{M}_{x} = \begin{cases} \mathbf{V} = \{X, Y\} \\ \mathbf{U} = \{U_{XY}, U_{X}, U_{Y}\} \\ \mathscr{F} = \begin{cases} \mathbf{X} = \mathbf{x} \\ Y = f_{Y}(\mathbf{x}, U_{Y}, U_{XY}) \\ P(\mathbf{U}) \end{cases}$$

Interventional Distribution

$$P(\mathbf{V} | \frac{do(X = x)}{do(X = x)}) \doteq P_{\mathcal{M}_x}(\mathbf{V})$$

Can we *predict* better the value of Y after making an intervention do(X = x)?

 $\exists x \text{ s.t. } P_{\mathcal{M}_x}(Y=y) \neq P(Y=y) \implies X \text{ is } a \text{ cause of } Y_{19}$



Structural Equation Model (SEM)

$$\mathcal{M} = \begin{cases} \mathbf{V} = \{X, Y, Z\} \\ \mathbf{U} = \{\epsilon_X, \epsilon_Y, \epsilon_Z\} \\ \mathcal{F} = \begin{cases} Z = \beta_{Z0} + \epsilon_Z \\ X = \beta_{X0} + \beta_{XZ} Z + \epsilon_X \\ Y = \beta_{Y0} + \beta_{YZ} Z + \beta_{YX} X + \epsilon_Y) \\ \mathbf{U} \sim \mathcal{N} \begin{pmatrix} \mathbf{0}, \Sigma = \begin{bmatrix} \sigma_X & 0 & 0 \\ 0 & \sigma_Y & 0 \\ 0 & 0 & \sigma_Z \end{bmatrix} \end{pmatrix} \end{cases}$$

- Linear functions \bullet
- Normal distribution \bullet
- Markovianity / Causal Sufficiency: Error terms in U are independent of each other (diagonal covariance matrix).

Full specification of an SCM requires parametric and distributional assumptions. Estimation of such models usually requires strong assumptions (e.g., Markovianity).



SCM: Encoder of <u>Functional</u> Knowledge



The knowledge required to fully specify an SCM is usually *unavailable* in practice.

Is it possible to identify the effect of interventions from observational data without fully specifying the SCM (i.e., in a non-parametric fashion)?

Yes, with structural knowledge encoded as a causal diagram!



Encoding Structural Causal Knowledge Acyclic Directed Acyclic Graph (ADMG) Causal Diagrams



Structural Causal Model (SCM) $\mathcal{M} = \langle \mathbf{V}, \mathbf{U}, \mathcal{F}, P(\mathbf{u}) \rangle$

$$\mathcal{M} = \begin{cases} \mathbf{V} = \{A, B, C\} \\ \mathbf{U} = \{U_A, U_B, U_C, U_{AB}\} \\ \\ \mathcal{M} = \begin{cases} A \leftarrow f_A(U_{AB}, U_A) \\ B \leftarrow f_B(U_{AB}, U_B) \\ C \leftarrow f_C(A, B, U_C) \end{cases} \\ P(\mathbf{U}) \end{cases}$$

An SCM $\mathcal{M} = \langle \mathbf{V}, \mathbf{U}, \mathcal{F}, P(\mathbf{u}) \rangle$ induces a causal diagram such that, for every $V_i, V_j \in \mathbf{V}$: $V_i \rightarrow V_j$, if V_i appears as argument of $f_i \in \mathscr{F}$.

Induced Causal Diagram







Structural Causal Model (SCM) $\mathcal{M} = \langle \mathbf{V}, \mathbf{U}, \mathcal{F}, P(\mathbf{u}) \rangle$

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 $V_i \leftrightarrow V_j$ if the corresponding $U_i, U_i \in \mathbf{U}$ are correlated or f_i, f_j share some argument $U \in \mathbf{U}$.





Structural Causal Model (SCM) $\mathcal{M} = \langle \mathbf{V}, \mathbf{U}, \mathcal{F}, P(\mathbf{u}) \rangle$

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D-Separation and Implied Conditional Independencies

1. Is a non-collider and is in \mathbb{Z} ; or

2. Is a collider and neither it nor any of its descendants in \mathbb{Z} .

a set of variables \mathbb{Z} if and only if p contains an inactive triplet in it.



Global Markov property: $(X \perp Y \mid Z)_G \Rightarrow (X \perp Y \mid Z)_P$

- **Definition (inactive):** A triplet $\langle V_i, V_m, V_j \rangle$ is said to be *inactive* relative to a set \mathbb{Z} if the middle node V_m :
- **Definition (d-separation):** A path p in a causal diagram G is said to be *d-separated* (or blocked) by
- A set Z d-separates X and Y if and only if Z blocks every path between a node in X and a node in Y.
 - Does \mathbb{Z} d-separates X and Y? **Z:** ∇ {*B*} {*W*} {*W*} {*B*, *W*}
 - We have that $(X \perp Y)_G$, $(X \perp Y \mid B)_G$, and $(X \perp Y \mid W)_G$, but $(X \perp Y \mid B, W)_G$

D-separations in G imply conditional independencies in P









Graphically Explaining Causes and Predictors



Markov Blanket (MB) of *V*: the bidirected connected component (district) of *V* (excluding *V* itself) and the parents of the district of *V*, i.e.: $mb_G(V) = dis_G(V) \cup Pa_G(dis_G(V)) \setminus \{V\}$



Randomized Experiments

Randomized Experiments / Control Trials (e.g. RCT) allow the identification of causal effects by leveraging randomization of the treatment assignment.







Pearl's Inferential Hierarchy

Associational vs Interventional vs Counterfactual



What is induced by the SCM?



















Data

usal DiagramsPotential SCMs
$$\mathcal{M}_{11} = \langle \mathbf{V}, \mathbf{U}_1, \mathcal{F}_{11}, P_{11}(\mathbf{u}_1) \rangle$$
: $\mathcal{M}_{1k_1} = \langle \mathbf{V}, \mathbf{U}_1, \mathcal{F}_{1k_1}, P_{1k_1}(\mathbf{u}_1) \rangle$ True $\mathcal{M}_{21} = \langle \mathbf{V}, \mathbf{U}_2, \mathcal{F}_{21}, P_{21}(\mathbf{u}_2) \rangle$: $\mathcal{M}_{2k_2} = \langle \mathbf{V}, \mathbf{U}_2, \mathcal{F}_{2k_2}, P_{2k_2}(\mathbf{u}_2) \rangle$: $\mathcal{M}_{31} = \langle \mathbf{V}, \mathbf{U}_3, \mathcal{F}_{31}, P_{31}(\mathbf{u}_3) \rangle$: $\mathcal{M}_{3k_3} = \langle \mathbf{V}, \mathbf{U}_3, \mathcal{F}_{3k_3}, P_{3k_3}(\mathbf{u}_3) \rangle$: $\mathcal{M}_{41} = \langle \mathbf{V}, \mathbf{U}_4, \mathcal{F}_{4k_4}, P_{4k_4}(\mathbf{u}_4) \rangle$: $\mathcal{M}_{51} = \langle \mathbf{V}, \mathbf{U}_5, \mathcal{F}_{51}, P_{51}(\mathbf{u}_5) \rangle$: $\mathcal{M}_{5k_5} = \langle \mathbf{V}, \mathbf{U}_5, \mathcal{F}_{5k_5}, P_{5k_5}(\mathbf{u}_5) \rangle$:

Encoded Knowledge / Assumptions











Ladder of Causation



Fask / Language	Typical Question	Examples
Structural Causal Model	What if I had acted differently?	Was it the aspirir that stopped my headache?
- Reinforcement ausal Bayes Net)	What if I do X? What would Y be if I intervene on X?	Will my headache be cured if I take aspirin?
L- (Un)Supervised Decision trees, Deep nets,)	What if I see? How would seeing X change my belief in Y?	What does a symptom tell us about the disease?

* Book of Why & On Pearl's Hierarchy and the Foundations of Causal Inference, E. Bareinboim, J. Correa, D. Ibeling, T. Icard, in press. <u>https://causalai.net/r60.pdf</u>


Ladder of Causation



ask / Language	Typical Question	Examples
layer infer	cences:	Was it the aspirin that stopped my headache?

most of the inferences are about causal effects (policies, treatments, decisions)

Vill my headache e cured if I take aspirin?

(most of the available data is observational, passively collected

What does a symptom tell us about the disease?

* Book of Why & On Pearl's Hierarchy and the Foundations of Causal Inference, E. Bareinboim, J. Correa, D. Ibeling, T. Icard, in press. <u>https://causalai.net/r60.pdf</u>



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 $\Im K_5$

 \mathcal{L}_{2}







Association vs Causation



Will we be able to decide the true relationship just by "seeing" more data?



https://xkcd.com/925/ - Creative Commons Attribution-NonCommercial 2.5 License.





Causal Effect Identification

Graphical Criteria, Do-Calculus, and ID-Algorithm



Causal Effect

Examples:

- Average Treatment Effect (ATE) for discrete treatments: $\mathbb{E}[\mathbf{Y} | do(\mathbf{X} = \mathbf{x}')] - \mathbb{E}[\mathbf{Y} | do(\mathbf{X} = \mathbf{x}')$ defined for two treatment levels \mathbf{x}' and \mathbf{x} of \mathbf{X} .
- Average Treatment Effect (ATE) for continuous treatments, $\partial \mathbb{E}[Y_i | do(X_j = x_j)]$, for all $Y_i \in \mathbf{Y}$, and $X_i \in \mathbf{X}$. ∂x_i The derivative shows the rate of change of Y w.r.t. do(X = x)

The causal effect of a (set of) treatment variable(s) X on a (set of) outcome variable(s) Y is a quantity derived from $P(\mathbf{Y} | do(\mathbf{X}))$ that tells us how much \mathbf{Y} changes due to an intervention $do(\mathbf{X} = \mathbf{x})$.

$$\mathbf{X}$$
 where $\mathbb{E}[\mathbf{Y} | do(\mathbf{X} = \mathbf{x})] = \sum_{\mathbf{y} \in \Omega_{\mathbf{Y}}} \mathbf{y} P(\mathbf{y} | do(\mathbf{x} = \mathbf{x}))$

Jacobian of
$$\mathbb{E}[\mathbf{Y} | do(\mathbf{X} = \mathbf{x})]$$
, where
 $\mathbb{E}[\mathbf{Y} | do(\mathbf{X} = \mathbf{x})] = \int_{\Omega_y} \mathbf{y} P(\mathbf{y} | do(\mathbf{x})) d\mathbf{y}$

and $\Omega_{\mathbf{V}}$ is the space of all possible values that **Y** might take on



Classical Causal Effect Identification



Conference on Artificial Intelligence (AAAI 2002), pp. 567–573, Menlo Park, CA, 2002. AAAI Press/MIT Press.

• Tian, J. and Pearl, J. A General Identification Condition for Causal Effects. In Proceedings of the Eighteenth National



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The Effect Identification Problem

Causal Effect Identifiability: The causal effect of a (set of) treatment variable(s) X on a (set of) outcome variable(s) \mathbf{Y} is said to be identifiable from a causal diagram G and the probability of the observed variables $P(\mathbf{V})$ if the interventional distribution $P(\mathbf{Y} | do(\mathbf{X}))$ is *uniquely computable*, i.e., if for every pair of SCMs \mathcal{M}_1 and \mathcal{M}_2 that induce G and $P^{\mathcal{M}_1}(\mathbf{V}) = P^{\mathcal{M}_2}(\mathbf{V}) = P(\mathbf{V}) > 0$, $P^{\mathcal{M}_1}(\mathbf{Y} \mid do(\mathbf{X})) = P^{\mathcal{M}_2}(\mathbf{Y} \mid do(\mathbf{X})) = P(\mathbf{Y} \mid do(\mathbf{X}))$.



In words, causal effect identifiability means that, no matter the form of true SCM, for all models \mathcal{M} agreeing with $\langle G, P(\mathbf{V}) \rangle$, they also agree in $P(\mathbf{y} | do(\mathbf{x}))$.





The Effect Identification Problem

Causal Effect Identifiability: The causal effect of a (set of) treatment variable(s) X on a (set of) outcome variable(s) \mathbf{Y} is said to be identifiable from a causal diagram G and the probability of the observed variables $P(\mathbf{V})$ if the interventional distribution $P(\mathbf{Y} | do(\mathbf{X}))$ is *uniquely computable*, i.e., if for every pair of SCMs \mathcal{M}_1 and \mathcal{M}_2 that induce G and $P^{\mathcal{M}_1}(\mathbf{V}) = P^{\mathcal{M}_2}(\mathbf{V}) = P(\mathbf{V}) > 0$, $P^{\mathcal{M}_1}(\mathbf{Y} \mid do(\mathbf{X})) = P^{\mathcal{M}_2}(\mathbf{Y} \mid do(\mathbf{X})) = P(\mathbf{Y} \mid do(\mathbf{X}))$.



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Tools for Causal Identification



Pearl, J. (2000). Causality: Models, Reasoning, and Inference. Cambridge University Press, New York. http:// dx.doi.org/10.1017/CBO9780511803161

Jin Tian. Studies in causal reasoning and learning. PhD thesis, University of California, Los Angeles, 2002.





interventional distributions $P_{\mathbf{x}}(\mathbf{V})$, for any $\mathbf{X} \subseteq \mathbf{V}$. It follows that $P_{\mathbf{x}}(\mathbf{V})$ factorizes as:

$$P_{\mathbf{x}}(\mathbf{v}) \doteq P(\mathbf{v} | do(\mathbf{x})) = \prod_{V_i \in \mathbf{V} \setminus \mathbf{X}} P_{\mathbf{x}}(v_i | pa_i) \Big|_{\mathbf{X} = \mathbf{x}}$$
$$= \prod_{V_i \in \mathbf{V} \setminus \mathbf{X}} P(v_i | pa_i) \Big|_{\mathbf{X} = \mathbf{x}}$$

Causal Effect of X on Y: $P(\mathbf{y} | do(\mathbf{x})) =$

- In Markovian Models, the joint interventional distribution (and hence any causal effect) is always identifiable.

Truncated Factorization – Markovian: Let G be a causal diagram for the collection \mathbf{P}_* of all

Follows from $P_{\mathbf{x}}(\mathbf{v}) \doteq P(\mathbf{v} | do(\mathbf{x}))$ being *Markov* relative to $G_{\overline{\mathbf{X}}}$

Markovian SCMs have the modularity property, i.e., $P_{\mathbf{x}}(v_i | pa_i) = P(v_i | pa_i)$

$$\sum_{\mathbf{V}\setminus(\mathbf{Y}\cup\mathbf{X})}\prod_{V_i\in\mathbf{V}\setminus\mathbf{X}}P(v_i|pa_i)\Big|_{\mathbf{X}=\mathbf{X}}$$

• This factorization is a.k.a "manipulation theorem" (Spirtes et al. 1993) or G-computation (Robins 1986, p. 1423).



Example: Identifiable Effect

Causal Effect of X on Y:



 $P(x, y, z) = P(z)P(x \mid z)P(y \mid x, z)$



 $P(\mathbf{y} | do(\mathbf{x})) = \sum_{\mathbf{V} \setminus (\mathbf{Y} \cup \mathbf{X})} \prod_{V_i \in \mathbf{V} \setminus \mathbf{X}} P_{\mathbf{x}}(v_i | pa_i) \Big|_{\mathbf{X} = \mathbf{x}}$ $G_{\overline{X}}$ do(X = x)

 $P(y, z \mid do(x)) = P(z)P(y \mid x, z)$

 $\implies P(y \mid do(x)) = \sum P(z)P(y \mid x, z)$ \mathcal{Z}



Adjustment over parents:

Let G be a causal graph with **no unmeasured parents**. Then, the effect of \mathbf{X} on \mathbf{Y} is given by: $P(\mathbf{y} | do(\mathbf{x})) = \sum P\left(\mathbf{y} | \mathbf{x}, \mathbf{pa}_{\mathbf{x}}\right) P\left(\mathbf{pa}_{\mathbf{x}}\right)$ pa_v



 $Pa_x = \{Z_1, Z_2\}$

Proof follows from the truncated factorization for Markovian models!

$$\mathbf{X} = \{X\}$$
$$\mathbf{Y} = \{Y\}$$
$$\mathbf{Pa_X} = \{Z_1, Z_2\}$$

 z_1, z_2

 $P(y \mid do(x)) = \sum_{x \in A} P(y \mid do(x)) = \sum_{x$



Adjustment over parents:

Let G be a causal graph with **no unmeasured parents**. Then, the effect of \mathbf{X} on \mathbf{Y} is given by: $P(\mathbf{y} | do(\mathbf{x})) = \sum P\left(\mathbf{y} | \mathbf{x}, \mathbf{pa}_{\mathbf{x}}\right) P\left(\mathbf{pa}_{\mathbf{x}}\right)$ pa_v



 $Pa_{r} = \{Z_{1}, Z_{2}\}$

 $P(\mathbf{v}|)$

Proof follows from the truncated factorization for Markovian models!

$$X = \{X\}$$

 $Y = \{Y\}$
 $Pa_X = \{Z_1, Z_2\}$

$$do(x)) = \sum_{z_1, z_2} P(y | x, z_1, z_2) P(z_1, z_2)$$

After conditioning on the parents, the association between X and Y is only due to the direct path.



Adjustment over parents:

Let G be a causal graph with **no unmeasured parents**. Then, the effect of \mathbf{X} on \mathbf{Y} is given by: $P(\mathbf{y} | do(\mathbf{x})) = \sum P\left(\mathbf{y} | \mathbf{x}, \mathbf{pa}_{\mathbf{x}}\right) P\left(\mathbf{pa}_{\mathbf{x}}\right)$ pa_v



 $Pa_{x} = \{Z_{2}\}$ $U_x = \{U_{X,Z2}\}$

P(y | do(x)) = ?





Adjustment over parents:

Let G be a causal graph with **no unmeasured parents**. Then, the effect of \mathbf{X} on \mathbf{Y} is given by: $P(\mathbf{y} | do(\mathbf{x})) = \sum P\left(\mathbf{y} | \mathbf{x}, \mathbf{pa}_{\mathbf{x}}\right) P\left(\mathbf{pa}_{\mathbf{x}}\right)$ pa



 $Pa_x = \{Z_2\}$ $U_x = \{U_{X,Z2}\}$ $P(y | do(x)) = \sum P(y | x, z_1, z_2) P(z_1, z_2)$ z_1, z_2





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$$P(y | do(x)) = \sum_{z_1, z_2} P(y | x, z_1, z_2) P(z_1, z_2)$$

After conditioning on the $\{Z_1, Z_2\}$, the association between X and Y is also due to a spurious / confounding path.



Backdoor Adjustment

Let X be a set of treatment variables and Y a set of outcome variables in the causal graph G. If there exists a set Z such that:

- 2. no node in \mathbb{Z} is a descendant of a variable $X \in \mathbb{X}$ (all variables in \mathbb{Z} are pre-treatment)

Then, Z satisfies the *backdoor criterion* for (X, Y) and, then the effect of X on Y is given by:

Z, a set of covariates, admissible for backdoor adjustment

Judea Pearl. Comment: Graphical models, causality and intervention. Stat. Sci., 8:266–269, 1993.

Also known as *confounding paths*, or backdoor paths.

1. for every $X \in \mathbf{X}$ and $Y \in \mathbf{Y}$, \mathbf{Z} blocks every path between X and Y that has an arrow into X, and







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Estimation via Propensity Scores

Theorem: If the set \mathbb{Z} satisfies the parent / backdoor criterion w.r.t. the ordered pair (X, Y) in the causal graph G, then the causal effect of X on Y is identifiable (uniquely computable) and given by:

Backdoor Adjustment \equiv **Conditional Ignorability:** $Y_{r} \perp X \mid \mathbb{Z}$

> Only if \mathbb{Z} satisfies the BD criterion, Inverse Probability Weighting/ Propensity Score can be used to estimate P(y | do(x)).

 $P(y \mid do(x)) = \sum P(y \mid x, \mathbf{z})P(\mathbf{z})$ $= \sum \frac{P(y \mid x, z) P(x \mid z) P(z)}{P(x \mid z)}$ $Z = \{Z_1\}$ $\mathbf{Z} = \{Z_1, Z_3\}$ $\frac{P(y, x, z)}{P(x \mid z)}$ propensity score neural nets Ζ





What if backdoor adjustment does not work?

Identification via Front-Door Adjustment

Let X be a set of treatment variables and Y a set of outcome variables in the causal graph G. If there exists a set M such that:

- 1. M intercepts all directed paths from any vertex $X \in \mathbf{X}$ to any vertex $Y \in \mathbf{Y}$;
- 2. There is no unblocked back-door path from any vertex $X \in \mathbf{X}$ to vertex $M \in \mathbf{M}$; and
- 3. All back-door paths from any vertex $M \in \mathbf{M}$ to any vertex $Y \in \mathbf{Y}$ are blocked by \mathbf{X} .

Then, \mathbf{M} satisfies the *front-door criterion* and, then the effect of \mathbf{X} on \mathbf{Y} is given by:



$$P(\mathbf{m} | \mathbf{x}) \sum_{\mathbf{x}'} P(\mathbf{y} | \mathbf{m}, \mathbf{x}') P(\mathbf{x}')$$

Many scenarios beyond back-door and front-door!





Conditional Front-Door

$$P(y | do(x)) = \sum_{m,z} P(m | x, z) \qquad P(y | do(x)) = -\frac{4}{2}$$
$$\sum_{x'} P(y | m, x', z) P(x', z)$$



Napkin

Unnamed

$$\sum_{z_2} P(x, y | z_1, z_2) P(z_2)$$

$$\sum_{z_2} P(x | z_1, z_2) P(z_2)$$

$$P(y | do(x)) = \sum_{z_2, z_3} P(y | x, z_1, z_2, z_3)$$
$$\sum_{z_1} P(z_3 | x, z_1) P(z_1)$$

And many others....





Do-Calculus (a.k.a. Causal Calculus)

Theorem: Let X, Y, Z, W be any disjoint subjects of variables. **Rule 1** (Insertion/Deletion of Observations) $P(\mathbf{y} | do(\mathbf{w}), \mathbf{x}, \mathbf{z}) = P(\mathbf{y} | do(\mathbf{w}), \mathbf{z}), \text{ if } (\mathbf{Y} \perp \mathbf{X} | \mathbf{Z}, \mathbf{W})_{G_{\overline{\mathbf{w}}}}$ **Rule 2** (Exchange of Actions and Observations) $P(\mathbf{y} | do(\mathbf{w}), \frac{do(\mathbf{x})}{do(\mathbf{x})}, \mathbf{z}) = P(\mathbf{y} | do(\mathbf{w}), \mathbf{x}, \mathbf{z}), \text{ if } (\mathbf{Y} \perp \mathbf{X} | \mathbf{Z}, \mathbf{W})_{G_{\overline{W}X}}$ **Rule 3** (Insertion/Deletion of Actions) $P(\mathbf{y} | do(\mathbf{w}), do(\mathbf{x}), \mathbf{z}) = P(\mathbf{y} | do(\mathbf{w}), \mathbf{z}), \text{ if } (\mathbf{Y} \perp \mathbf{X} | \mathbf{Z}, \mathbf{W})_{G_{\overline{\mathbf{W}}, \overline{\mathbf{X}(\mathbf{Z})}}}$

X(Z): set of X-nodes that are not ancestors of any Z-node in $G_{\overline{W}}$.

- Graphical conditions implying invariances between observational (\mathscr{L}_1) and interventional (\mathscr{L}_2) distributions

 $G_{\overline{W}X}$: graph G after removing the incoming arrows into W and the outgoing arrows from X;





X

$$P(y | do(x)) = \sum_{m}^{m} P(y | do(x), m) P(m | do(x))$$

= $\sum_{m}^{m} P(y | do(x), do(m)) P(m | do(x))$
= $\sum_{m}^{m} P(y | do(x), do(m)) P(m | x)$
= $\sum_{m}^{m} P(y | do(m)) P(m | x)$
= $\sum_{x'}^{m} \sum_{m}^{m} P(y | do(m), x') P(x' | do(m)) P(m | x)$
= $\sum_{x'}^{m} \sum_{m}^{m} P(y | m, x') P(x' | do(m)) P(m | x)$



Probability Axioms

Rule 2

Rule 2

Rule 3

Probability Axioms

Rule 2



Rule 3



The Identify (ID) Algorithm

Algorithm 1 ID(x, y) given Causal Diagram G

Input: two disjoint sets $\mathbf{X}, \mathbf{Y} \subset \mathbf{V}$ Output: Expression for $P_{\mathbf{x}}(\mathbf{y})$ or FAIL

- 1: Let $\mathbf{D} = \operatorname{An}(\mathbf{Y})_{\mathcal{G}_{\mathbf{V}\setminus\mathbf{X}}}$
- 2: Let the c-components of $\mathcal{G}_{\mathbf{D}}$ be $\mathbf{D}_i, i = 1, \dots, k$
- 3: $P_{\mathbf{x}}(\mathbf{y}) = \sum_{\mathbf{d} \setminus \mathbf{y}} \prod_{i} \text{IDENTIFY}(\mathbf{D}_{i}, \mathbf{V}, P)$
- 4: function IDENTIFY($\mathbf{C}, \mathbf{T}, Q = Q[\mathbf{T}]$)
- 5: if C = T then return Q[T]

/* Let S^B denote the c-component of $\{B\}$ in $\mathcal{G}_{\mathbf{T}}$ */

- 6: if $\exists B \in \mathbf{T} \setminus \mathbf{C}$ such that $S^B \cap Ch(B) = \emptyset$ then
- 7: Compute $Q[\mathbf{T} \setminus \{B\}]$ from Q; \triangleright Lemma 1
- 8: **return** IDENTIFY($\mathbf{C}, \mathbf{T} \setminus \{B\}, Q[\mathbf{T} \setminus \{B\}]$)
- 9: **else**
- 10: **throw** FAIL

Jin Tian. Studies in causal reasoning and learning. PhD thesis, University of California, Los Angeles, 2002. Link

Lemma 1. Given a causal diagram \mathcal{D} over $\mathbf{V}, X \in \mathbf{T} \subseteq \mathbf{V}$, and $P_{\mathbf{v}\setminus \mathbf{t}}$, i.e., an expression for $Q[\mathbf{T}]$. If X is not in the same c-component with a child in $\mathcal{D}_{\mathbf{T}}$, then $Q[\mathbf{T} \setminus {X}]$ is identifiable and given by

$$Q[\mathbf{T} \setminus \{X\}] = \frac{P_{\mathbf{v} \setminus \mathbf{t}}}{Q[S^X]} \times \sum_x Q[S^X]$$
(2)

where S^X is the c-component of X in $\mathcal{D}_{\mathbf{T}}$ and $Q[S^X]$ is computable from $P_{\mathbf{v}\setminus\mathbf{t}}$ by [Tian, 2002, Lemma 11].

n nma 1 ·])





Causal Effect Identification



Conference on Artificial Intelligence (AAAI 2002), pp. 567–573, Menlo Park, CA, 2002. AAAI Press/MIT Press.

• Tian, J. and Pearl, J. A General Identification Condition for Causal Effects. In Proceedings of the Eighteenth National





More on Causal Effect Identification

Identification from observational and experimental data:

Lee, S., Correa, J., and Bareinboim, E. (**2019**). General identifiability with arbitrary surrogate experiments. In *Proceedings of the 35th Conference on Uncertainty in Artificial Intelligence*, volume 35, Tel Aviv, Israel. AUAI Press. Link

Identification of stochastic/soft (and possibly imperfect) interventions:

Correa, J. and Bareinboim, E. (2020). A calculus for stochastic interventions: Causal effect identification and surrogate experiments. In *Proceedings of the 34th AAAI Conference on Artificial Intelligence*, New York, NY. AAAI Press. Link

Identification and Estimation via Deep Neural Networks:

Xia, K., Lee, K.-Z., Bengio, Y., and Bareinboim, E. (2021). The causal-neural connection: Expressiveness, learnability, and inference. *Advances in Neural Information Processing Systems*, 34. Link





Identification and Estimation via Deep Neural Networks

True Model: SCM M*

> Trained Model: G-NCM \widehat{M}

G

X

V: Endogenous variables U: Create one for every bidirected clique \mathcal{F} : Feedforward neural network for each variable in V with parents from the graph $P(\widehat{\mathbf{U}})$: All Unif(0,1)



Inductive bias based on the causal diagram: the enforced constraints empower the NCM with the ability to solve causal inference tasks.





Expressiveness of NCMs



Thm: For any SCM \mathcal{M}^* , there exists an NCM \widehat{M} such that \widehat{M} matches \mathcal{M}^* on all three PCH layers! This does not imply that the estimated NCM \widehat{M} matches the true SCM $\mathscr{M}^*!$





Solution: A Neural Algorithm for Identification

Algorithm 1: Identifying/estimating queries with NCMs.

Input : causal query $Q = P(\mathbf{y} \mid do(\mathbf{x})), L_1$ data $P(\mathbf{v})$, and causal diagram \mathcal{G} **Output :** $P^{\mathcal{M}^*}(\mathbf{y} \mid do(\mathbf{x}))$ if identifiable, FAIL otherwise. 1 $\widehat{M} \leftarrow \text{NCM}(\mathbf{V}, \mathcal{G})$ 2 $\boldsymbol{\theta}_{\min}^* \leftarrow \arg\min_{\boldsymbol{\theta}} P^{\widehat{M}(\boldsymbol{\theta})}(\mathbf{y} | do(\mathbf{x})) \text{ s.t. } L_1(\widehat{M}(\boldsymbol{\theta})) = P(\mathbf{v}) \checkmark$ 3 $\boldsymbol{\theta}_{\max}^* \leftarrow \arg \max_{\boldsymbol{\theta}} P^{\widehat{M}(\boldsymbol{\theta})}(\mathbf{y} | do(\mathbf{x})) \text{ s.t. } L_1(\widehat{M}(\boldsymbol{\theta})) = P(\mathbf{v})$ 4 if $P^{\widehat{M}(\boldsymbol{\theta}_{\min}^*)}(\mathbf{y} \mid do(\mathbf{x})) \neq P^{\widehat{M}(\boldsymbol{\theta}_{\max}^*)}(\mathbf{y} \mid do(\mathbf{x}))$ then return FAIL 6 else return $P^{\widehat{M}(\boldsymbol{\theta}_{\min}^*)}(\mathbf{y} \mid do(\mathbf{x}))$ // choose min or max 7 arbitrarily

The approach is equivalent to established symbolic approaches (Thm. 4), and in identifiable cases, the result is an NCM that can serve as a proxy model for estimating the query (Corol. 2).



Maximize and minimize the induced causal query Q while maintaining L_1 -consistency (can be done with likelihood estimation).

> **Thm.** : Q is identifiable if and only if they match!

Corol: If Q is identifiable, then we can compute it by performing the mutilation procedure on M!





Can we relax some causal assumptions?



Causal Effect Identification



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• Tian, J. and Pearl, J. A General Identification Condition for Causal Effects. In Proceedings of the Eighteenth National





Is a Causal Diagram Still Too Much?

Question:

- Causal diagrams are powerful tools that allow for inferences based on weaker knowledge (structural invariances) than the encoded in the true, underlying SCM.
 - Still, structural knowledge for every pair of variables may not be available in many real-world, complex, high-dimensional systems.

Is it possible to relax the assumption of having a fully specified causal diagram and still be able to identify a causal effect?



Partially Understood Systems

A) Age
(B) Blood pressure
(C) Comorbidities
(D) Medication history
(X) Lisinopril
(S) Sleep Quality
(Y) Stroke





A causal diagram cannot be specified given the existing knowledge!

How can we identify P(y | do(x)) in this case?



Cluster DAGs (C-DAGs)

A) Age (*B*) Blood pressure (*C*) Comorbidities (D) Medication history (X) Lisinopril (S) Sleep Quality (*Y*) Stroke



A cluster DAG $G_{\mathbf{C}}$ over a given partition $\mathbf{C} = \{\mathbf{C}_1, \dots, \mathbf{C}_k\}$ of V is compatible with a causal diagram G over V if for every $C_i, C_i \in C$:

•
$$\mathbf{C}_i \to \mathbf{C}_j$$
 if $\exists V_i \in \mathbf{C}_i$ and $V_j \in \mathbf{C}_j$ such

• $\mathbf{C}_i \leftrightarrow \mathbf{C}_i \text{ if } \exists V_i \in \mathbf{C}_i \text{ and } V_i \in \mathbf{C}_i \text{ such that } V_i \leftrightarrow V_j$

and $G_{\mathbf{C}}$ contains no cycles.

$\{\{X\}, \{S\}, \{Y\}, \{A, B, C, D\}\}$

- that $V_i \rightarrow V_i$



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Partially Understood Systems

Many causal diagrams are compatible with the current knowledge!



Can be seen as an *equivalence class* of causal diagrams, where any relationships are allowed among the variables within each cluster.

Can we infer causal effects without deciding on any one particular causal diagram?




C-DAG: Flexible Encoder of Model Assumptions





N clusters of size one (full knowledge - DAG)

(partial knowledge - C-DAG)





One cluster of size N (no knowledge)

low

trivial



C-DAG: Flexible Encoder of Model Assumptions



N clusters of size one (full knowledge - DAG)

Clusters are manually created by domain experts:

- to communicate relationships among semantically meaningful entities.





(partial knowledge - C-DAG)

One cluster of size N (no knowledge)

- due to lack of knowledge, consensus, or interest on the internal causal structure;



Identification of Causal Effects from C-DAGs



Anand, T. V., Ribeiro A. H., Tian, J., & Bareinboim, E. (2023). Causal Effect Identification in Cluster DAGs. In Proceedings of the Thirty-Seventh AAAI Conference on Artificial Intelligence.





Effect Identifiabiliy given a C-DAG



An effect identifiable in a C-DAG G_{C} is identifiable in all compatible causal diagrams G using the same identification formula!



Effect Non-Identifiabiliy given a C-DAG



 $P(y \mid do(x))$ is not identifiable

An effect is not identifiable in a C-DAG $G_{\mathbf{C}}$ if there exists at least one compatible causal diagrams G in which the effect is not identifiable.







Beyond Backdoor Adjustment



Again, an effect identifiable in a C-DAG $G_{\mathbf{C}}$ is identifiable in all compatible causal diagrams G using the same identification formula!



What if no knowledge is available?

Causal Discovery:

In non-parametric settings, we can't learn the true causal diagram, but algorithms such as the Fast Causal Inference (FCI) can learn a graphical representation of its Markov equivalence class!

Zhang, J. (2008). On the completeness of orientation rules for causal discovery in the presence of latent confounders and selection bias. Artificial Intelligence, 172(16):1873–1896. Link

Can we learn a causal diagram \mathcal{G} from observational data?







Causal Discovery

Fast Causal Inference (FCI)

A constraint-based causal discovery algorithms that accounts for unobserved confounders



Causal Discovery

Goal: Learn a graphical representation of the Markov Equivalence Class from observational data.

Assumptions: the observed distribution is the marginal of a distribution P that satisfies the following conditions for the true causal diagram G (an **ADMG**):

- I-Map / Semi-Markov Condition: for any disjoint subsets X, Y and Z: 1) $(\mathbf{X} \perp \mathbf{Y} \mid \mathbf{Z})_G \Rightarrow (\mathbf{X} \perp \mathbf{Y} \mid \mathbf{Z})_P.$
- 2) Faithfulness Condition: for any disjoint subsets X, Y and Z: $(\mathbf{X} \perp \mathbf{Y} \mid \mathbf{Z})_P \Rightarrow (\mathbf{X} \perp \mathbf{Y} \mid \mathbf{Z})_G.$

Note: Estimation of the marginal distribution from limited data requires and **additional assumption**: 3) An adequate *conditional independence test* is available.

G is an *I-Map of* P *P* is **semi-Markov**

P is *faithful* to *G*





relative to G.



Fast Causal Inference (FCI) Algorithm

FCI: Learn a PAG \mathscr{P} representing the Markov Equivalence Class (MEC) from P, i.e.:

$(\mathbf{X} \perp \mathbf{Y} \mid \mathbf{Z})_{\mathcal{P}}$

Every non-circle edge mark represents an invariance in the MEC in terms of ancestral and non-ancestral relationships

Arrowhead \implies non-ancestrality Tail \implies ancestrally \implies non-invariance Circle

Zhang, J. (2008). On the completeness of orientation rules for causal discovery in the presence of latent confounders and selection bias. Artificial Intelligence, 172(16):1873–1896. Link

Jaber A., Ribeiro A. H., Zhang, J., Bareinboim, E. Causal Identification under Markov Equivalence - Calculus, Algorithm, and Completeness. In Proceedings of the 36th Annual Conference on Neural Information Processing Systems, NeurIPS 2022. (Link)

$$\Leftrightarrow (\mathbf{X} \perp \mathbf{Y} \mid \mathbf{Z})_{P;G}$$

Evaluated through m-separation

- $A \longrightarrow B \implies$ ancestrally
- $A \longrightarrow B \implies$ non-ancestrality
- $A \longleftrightarrow B \implies$ spurious association
- $A \longmapsto B \implies$ selection bias





Conditional Independence Tests

Gaussian errors and independent observations: partial correlation test

Fisher, R.A. (1921). On the "Probable Error" of a Coefficient of Correlation Deduced from a Small Sample. R package: <u>https://cran.r-project.org/web/packages/pcalg/</u>

Kernel-based non-parametric test:

Zhang, K., Peters, J., Janzing, D., & Schölkopf, B. (2012). Kernel-based conditional independence test and application in causal discovery. In: Uncertainty in artificial intelligence. AUAI Press; 2011. p.804–13 R package: <u>https://cran.r-project.org/web/packages/CondIndTests</u>

Continuous (conditional Gaussian) or Discrete (Binary, Ordinal, Multinomial) - Linear Regression

- data. Int J Data Sci Anal 6, 19–30. (Link)
- R package: <u>https://cran.r-project.org/web/packages/MXM/</u>

Gaussian errors and correlated observations (family data) :

Ribeiro A.H., Soler J.M.P. (2020). Learning Genetic and environmental graphical models from family data, Statistics in Medicine.

R package: <u>https://github.com/adele/FamilyBasedPGMs</u>

• Tsagris, M., Borboudakis, G., Lagani, V. et al. (2018) Constraint-based causal discovery with mixed





Learning Structural Invariances

 $\mathbf{V} = \{X, Y, Z\}$ $\mathbf{U} = \{U_x, U_Y, U_Z\}$ Conditional $\mathcal{M}_{1} = \begin{cases} \mathcal{K} \leftarrow f_{X}(U_{X}) \\ Z \leftarrow f_{Z}(X, Y, U_{Z}) \\ Y \leftarrow f_{Y}(U_{Y}) \end{cases}$ (in)dependencies Data $P(\mathbf{U})$ $P(\mathbf{v})$ $X \perp \!\!\!\!\perp Y$ $\int \mathbf{V} = \{X, Y, Z\}$ $\mathcal{M}_{N-1} = \begin{cases} \mathcal{V} = \{U_{XZ}, U_{YZ}, U_X, U_Y, U_Z\} \\ \mathcal{M}_{N-1} = \begin{cases} X \leftarrow f_X(U_{XZ}, U_X) \\ Z \leftarrow f_Z(Y, U_{XZ}, U_Z) \\ Y \leftarrow f_Y(U_Y) \end{cases}$ $X \not\perp Z$ $X \not\sqcup Y | Z$ $P(\mathbf{U})$ $\mathbf{V} = \{X, Y, Z\}$ $P(x, y, z) = P(z \mid x, y) P(x \mid y) P(y)$ $\mathbf{U} = \{U_{XZ}, U_{YZ}, U_X, U_Y, U_Z\}$ $\mathcal{M}_{N} = \begin{cases} X \leftarrow f_{X}(U_{XZ}, U_{X}) \\ Z \leftarrow f_{Z}(U_{XZ}, U_{YZ}, U_{Z}) \\ Y \leftarrow f_{Y}(U_{YZ}, U_{Y}) \end{cases}$ $= P(z \mid x, y) P(x) P(y)$ $P(\mathbf{U})$





Learning Structural Invariances



Zhang, J. (2008). On the completeness of orientation rules for causal discovery in the presence of latent confounders and selection bias. Artificial Intelligence, 172(16):1873–1896. Link





Other examples

Underlying Causal Diagram



Zhang, J. (2008). On the completeness of orientation rules for causal discovery in the presence of latent confounders and selection bias. *Artificial Intelligence*, 172(16):1873–1896. Link

Partial Ancestral Graph





Fast Causal Inference (FCI) Algorithm



observed in the data

Implied by the PAG $X \perp W$ using m-separation $X \perp Y \mid Z, W$

Z is not an ancestor of X or W.

Z and W are ancestors of Y.

Z is not confounded with Y.



PAG represents the Markov Equivalence Class



Partial Ancestral Graph (PAG)

Z is not an ancestor of X or W.

Z and W are ancestors of Y.

Z is not confounded with Y.



True (unknown) causal diagram

 $\begin{array}{c} X \coprod W \\ X \coprod Y | Z, W \end{array}$



More on Causal Discovery

Causal discovery from observational and experimental data:

- Learning and Reasoning, PMLR 177:253-274, 2022.
- Neural Information Processing Systems 2020.

• Gonçalo Rui Alves Faria, Andre Martins, Mario A. T. Figueiredo. Differentiable Causal Discovery Under Latent Interventions. Proceedings of the First Conference on Causal

• Kocaoglu, M., Jaber, A., Shanmugam, K., Bareinboim, E. Characterization and Learning of Causal Graphs with Latent Variables from Soft Interventions. In Proceedings of the 33rd Annual Conference on Neural Information Processing Systems. 2019.

• Jaber, A., Kocaoglu, M., Shanmugam, K., Bareinboim, E. Causal Discovery from Soft Interventions with Unknown Targets: Characterization & Learning. In Advances in



Causal Identification from PAGs

Effect Identification:

of marginal and conditional causal effect in PAGs!

Can we identify causal effects from the equivalence class?

- For Covariate Adjustment, we can use the Generalized Adjustment Criterion.
- Recently, we proposed complete calculus and algorithms for the identification
- Perkovic, E., Textor, J. C., Kalisch, M., & Maathuis, M. H. (2018). Complete graphical characterization and construction of adjustment sets in Markov equivalence classes of ancestral graphs. Journal of Machine Learning Research 18 (2018) 1-62
- Jaber A., Ribeiro A. H., Zhang, J., Bareinboim, E. (2022) Causal Identification under Markov Equivalence Calculus, Algorithm, and Completeness. In Proceedings of the 36th Annual Conference on Neural Information Processing Systems, NeurIPS. (Link)



General Identification in Markov Equivalence Classes



Jaber A., Ribeiro A. H., Zhang, J., Bareinboim, E. Causal Identification under Markov Equivalence - Calculus, Algorithm, and Completeness. In Proceedings of the 36th Annual Conference on Neural Information Processing Systems, NeurIPS 2022. (Link) 91

The CIDP and IDP algorithms are available at the PAGId R package: https://github.com/adele/PAGId

IDP / CIDP Solution yes/no $P(y \mid do(x)) = \sum P(y \mid x, z) P(z)$ Interventional Available Distribution Distributions









Effect Identifiabiliy given a PAG



$$P(y \mid do(x)) = \sum_{z} P(y \mid x, z) P(z)$$

An effect identifiable in a PAG \mathscr{P} is identifiable in all causal diagrams G in the Markov Equivalence Class using the same identification formula!





Effect Non-Identifiabiliy given a PAG



P(y | do(x)) is not identifiable

An effect not identifiable in a PAG \mathscr{P} is not identifiable in at least one causal diagrams G in the Markov Equivalence Class



X

 $P(y | do(x)) = \sum_{z} P(y | x, z) P(z)$

P(y | do(x)) is not identifiable



Causal Inference Workflow

Continuous Process of Scientific Discovery and Causal Hypothesis Refinement



new discoveries(t+1)



Causal Inference Workflow

Continuous Process of Scientific Discovery and Causal Hypothesis Refinement







There is much more to learn...

- 1. Generalized Effect Identification:
 - For when multiple observational and experimental datasets are available, possible under partial observability.
- 2. Partial Identification:
 - For when the effect is not point-identifiable, but an interval for it can be derived.
- 3. Effect Transportability:
 - For when the target is a different population/domain.
- 4. Counterfactual Identification:

- Identification of \mathscr{L}_3 quantities, such as $P(\mathbf{y}_{\mathbf{x}} | \mathbf{x}')$ and $P(\mathbf{y}_{\mathbf{x}} | \mathbf{x}', \mathbf{z})$.

- 5. Fairness Evaluation:
 - To identify path-specify effects related to protective variables.
- 6. Effect Estimation beyond backdoor scenarios:
 - Via doubly robust machine learning and different plug-in density estimators.



Causal inference can help overcome critical challenges in Artificial Intelligence, including robustness, generalizability, explainability, and fairness.

Causal Data Science: principled way of combining data and substantive knowledge about the phenomenon under investigation to generate causal explanations and better decision-making.

Recent developments for causal inference when knowledge is largely unavailable and coarse are expected to help the practice of causal data analysis and meet the growing demand in the Empirical Sciences for sound causal explanations and more robust and generalizable decision-making.





Thank you! :)

adele.ribeiro@uni-marburg.de

Feel free to reach out to me if you have any questions:

